

Minimax optimality of diffusion models

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- Let P^* be an arbitrary target measure. Set $X_0 \sim P^*$ and construct a diffusion model from n data D_n . Let \hat{Y}_T be the random element induced from the diffusion model after the sufficient iterations T . Then, what would be the worst case estimation rate, i.e.,

$$\sup_{X_0 \sim P^*} \mathbb{E}_{D_n} d(X_0, \hat{Y}_T) \lesssim n^{-\square}?$$

- How optimal the above rate is?
 - For \hat{P} any estimator of P^* , the following rate is called 'minimax optimal rate':

$$\inf_{\hat{P}} \sup_{P^*} \mathbb{E}_{D_n} d(P^*, \hat{P}) \gtrsim n^{-\square}.$$

- Can diffusion model achieve the minimax optimal rate? YES, at least nearly, when $d = TV$ or W_1 .
- I will focus on $d = W_1$; this one has more interesting intuition than TV .

Preliminaries

- Besov space: characterizes a function space with some ‘smoothness’.

- Fix the domain of X by $\Omega := [0, 1]^d$, a unit hypercube.
- For $f \in L^p(\Omega)$, define the ‘ r th-modulus of smoothness’:

$$w_{r,p}(f, t) := \sup_{\|h\| \leq t} \|\Delta_h^r(f)\|_{L^p(\Omega)},$$

where $\Delta_h^r(f)(x) := \sum_{j=0}^r \binom{r}{j} (-1)^{r-j} f(x + jh)$ if $x, x + h \in \Omega$, and 0 otherwise.

- E.g. $\Delta_h^1(f)(x) = f(x + h) - f(x)$; $\Delta_h^2(f)(x) = f(x + 2h) - 2f(x + h) + f(x)$.
- Let $r = \lfloor s \rfloor + 1$, and define a Besov semi-norm

$$|f|_{B_{p,q}^s(\Omega)} := \begin{cases} \left[\int_{\Omega} \left(\frac{w_{r,p}(f,t)}{t^s} \right)^q \frac{dt}{t} \right]^{\frac{1}{q}} & 0 < q < \infty, \\ \sup_{t>0} \frac{w_{r,p}(f,t)}{t^s} & q = \infty. \end{cases}$$

- If f satisfies $\|f\|_{B_{p,q}^s(\Omega)} := \|f\|_{L^p(\Omega)} + |f|_{B_{p,q}^s(\Omega)} < \infty$, then f is said to be in a Besov space $B_{p,q}^s(\Omega)$.
- Besov space is not a Banach space, but quasi-Banach space.
- Easier interpretaions by examples:
 - $B_{p,1}^s(\Omega) \hookrightarrow W_p^s(\Omega) \hookrightarrow B_{p,\infty}^s(\Omega)$. Particulary, $B_{2,2}^s(\Omega) = W_2^s(\Omega)$.
 - $B_{p,q}^s \approx W_p^s$, and q is just for some finer distinctions.
 - As in Sobolev embedding, $s > d/p$ implies the continuity of f .
 - Important example for later: $B_{\infty,1}^1(\Omega) \hookrightarrow Lip(\Omega) \hookrightarrow B_{\infty,\infty}^1$.

- Let $\Phi(L, W, S, B)$ be a L -layer W -width ReLU Deep neural network with the following structure:

$$\Phi(L, W, S, B)(x) = \left[\left(W^{(L)}(\cdot) + b^{(L)} \right) \circ \sigma \cdots \circ \sigma \left(W^{(1)}(\cdot) + b^{(1)} \right) \right](x). \quad (1)$$

- σ : ReLU activation function.
- L : Neural network depth.
- W : Neural network width, i.e., $W^{(l)} \in \mathbb{R}^{W \times W}$, $b^{(l)} \in \mathbb{R}^W$ for all $l = 1, \dots, L$.
- S : Sparsity parameter, i.e., $\sum_{l=1}^L \left[\|W^{(l)}\|_0 + \|b^{(l)}\|_0 \right] \leq S$.
- B : Norm constraint, i.e., $\max_{l=1, \dots, L} \left[\|W^{(l)}\|_\infty, \|b^{(l)}\|_\infty \right] \leq B$.

- Goal: Given only data $x_i \stackrel{i.i.d}{\sim} P^*$, generate more samples $x_i \sim P^*$.

- Procedure:

- 1 Assume P^* : initial distribution of some Ornstein–Uhlenbeck (OU) process, i.e., for $X_0 \sim P_0 = P^*$,

$$dX_t = -\beta_t X_t dt + \sqrt{2\beta_t} dB_t.$$

Note $X_t \rightarrow N(0, I)$ exponentially. We consider this process up to some timestep T .

- 2 Let $Y_0 \sim N(0, I)$. The goal is to construct a dynamical system Y_t s.t.

$$Y_t = X_{T-t} \Rightarrow Y_T = X_0 = X^*.$$

- 3 Then, the following SDE induces $Y_t = X_{T-t}$ (reverse process):

$$dY_t = \beta_{T-t}(Y_t + 2\nabla \log P_{T-t}(Y_t))dt + \sqrt{2\beta_{T-t}}dB_t.$$

- 4 Y_T is the distribution we desire, but we cannot obtain this as we do not know the P_t . Instead, assume for each t we have $\hat{s}(Y_t, t)$, an estimator of the score function $\nabla \log P_t(Y_t)$. Consider the estimator \hat{Y}_t :

$$d\hat{Y}_t = \beta_{T-t}(\hat{Y}_t + 2\hat{s}(\hat{Y}_t, T-t))dt + \sqrt{2\beta_{T-t}}dB_t.$$

- 5 To generate $x_i \sim P^*$, set $\hat{Y}_0^{(i)} \stackrel{i.i.d}{\sim} N(0, I) \Rightarrow x_i = \hat{Y}_T^{(i)} \approx P^*$, given \hat{s} is a *nice* estimator.

- To obtain the *nice* estimator $\widehat{s}(Y_t, T - t)$, one trains a function (typically DNN) with ‘score matching loss’. Fix some function class \mathcal{S} (typically DNN); then, for some fixed $\epsilon > 0$, define

$$\begin{aligned}\widehat{s} &= \operatorname{argmin}_{s \in \mathcal{S}} \frac{1}{n} \sum_{i=1}^n \int_{\epsilon}^T \mathbb{E}_{x_t \sim P_t(x_t | (x_0, i)_{i=1, \dots, n})} \left[\|s(x_t, t) - \nabla \log P_t(x_t | (x_0, i)_{i=1, \dots, n})\|^2 \right] dt \\ &\approx \operatorname{argmin}_{s \in \mathcal{S}} \mathbb{E}_{x_0 \sim P^*} \left[\int_0^T \mathbb{E}_{x_t \sim P_t(x_t | x_0)} \left[\|s(x_t, t) - \nabla \log P_t(x_t | x_0)\|^2 \right] dt \right].\end{aligned}$$

- Note we set the target as $\nabla \log P_t(x_t | x_0)$ instead of $\nabla \log P_t(x_t)$; this trick is sometimes called a score matching trick.

- Let \mathcal{P} to be a set of absolutely continuous probability measures on Ω with density f in a Besov space $B_{p,q}^s(\Omega)$ and bounded below and above by C_f^{-1} and C_f .

Theorem (Minimax optimality (NWB22))

For any estimator $\hat{P} \in \mathcal{P}$ constructed using n data $D_n = (x_i)_{i=1,\dots,n}$,

$$n^{-\frac{s+1}{2s+d}} \lesssim \inf_{\hat{P} \in \mathcal{P}} \sup_{P^* \in \mathcal{P}} \mathbb{E}_{D_n} W_1(P^*, \hat{P}).$$

Theorem (Diffusion models are nearly minimax optimal (OAS23))

For any $\delta > 0$, if we train the diffusion model with the score estimator $\hat{s}(x, t) \in \Phi(L, W, S, B)$ for some L, W, S, B that depends on n, d, p, s, T and $T \geq \frac{(s+1) \log n}{\min_t \beta_t(2s+d)}$, then

$$\sup_{X_0 \sim P^* \in \mathcal{P}} \mathbb{E}_{D_n} W_1(X_0, \hat{Y}_T) \lesssim n^{-\frac{s+1-\delta}{2s+d}}.$$

- DNN diffusion model is ‘nearly’ (the gap is $n^{\frac{\delta}{2s+d}}$) minimax optimal.
- In practice, the score loss blows up as $t \rightarrow 0$, so often one uses clipping $t \in [\epsilon, T]$ for some $\epsilon > 0$. The actual theorem is written w.r.t $\hat{Y}_{T-\epsilon}$, but for simplicity we assume $\epsilon = 0$.

- Result is from (NWB22)[Theorem 3, Proposition 3].
- Key idea: Observe the following calculation (MRCS10): for any $h \in C^1(\Omega)$,

$$\begin{aligned} \int_{\Omega} h(dP - dQ) &= \int_0^1 \frac{d}{dt} \left(\int h dP_t \right) dt = \int_0^1 \int_{\Omega} \nabla h \cdot v_t dP_t dt \\ &\leq \left(\int_0^1 \int_{\Omega} \|\nabla h\|^p dP_t dt \right)^{1/p} \left(\int_0^1 \int_{\Omega} \|v_t\|^q dP_t dt \right)^{1/q} \\ &\leq C^{1/p} \|\nabla h\|_{L^p(\Omega)} W_q(P, Q). \end{aligned}$$

- If $P, Q \in \mathcal{P}$, one can choose the optimal h in LHS to get $\|f_P - f_Q\|_{B_{1,\infty}^{-1}} \lesssim W_1(P, Q)$.
- Plug-in $P = P^*, Q = \hat{P}$, and $\|f_{P^*} - f_{\hat{P}}\|_{B_{1,\infty}^{-1}}$'s lower bound can be derived using the standard Besov space minimax estimation technique (KP92).

- To prove the minimax optimality of the DNN diffusion model, we first need the performance guarantee of DNN in general Besov function estimation.
- Consider the problem of estimating $f^* \in B_{p,q}^s(\Omega) \cap B_{L^\infty}(\Omega)(0, F)$ for some $F > 0$, with the data $y_i = f^*(x_i) + \epsilon_i$ with $\epsilon_i \stackrel{i.i.d}{\sim} N(0, \sigma^2)$ and $X \sim P$ where $\text{supp}(P) \subseteq \Omega$.

Theorem (DNN estimator of Besov function)

Let $\hat{f} := \operatorname{argmin}_{h \in \Phi(L, W, S, B)} \sum_{i=1}^n |y_i - h(x_i)|^2$ with L, W, S, B that depends on n, s, d, p . For all $f^* \in B_{p,q}^s(\Omega)(0, 1) \cap B_{L^\infty}(\Omega)(0, F)$ with some $F > 0$,

$$\mathbb{E}_{D_n} \left\| f^* - \hat{f} \right\|_{L^2(P)}^2 \lesssim n^{-\frac{2s}{2s+d}} (\log n)^3.$$

- The proof consists of two ingredients:
 - Approximation of Besov function by some DNN \tilde{f} (may depends on f^*).
 - Statistical learning theory to control the error between \hat{f} and any choice of the approximator \tilde{f} .
 - Total error is bounded by the above two errors $\left\| \hat{f} - \tilde{f} \right\|, \left\| \tilde{f} - f^* \right\|$.

- Optimal approximation error: For sufficiently large $N \in \mathbb{N}$, there exists L, W, S, B that depends on N, d, s, p s.t.

$$\sup_{f^* \in B_{B_{p,q}^s(\Omega)}(0,1)} \inf_{\tilde{f} \in \Phi(L, W, S, B)} \left\| \tilde{f} - f^* \right\| \lesssim N^{-\frac{s}{d}}.$$

- Basic strategy: two-stage approximation: $B_{p,q}^s(\Omega) \approx$ B-spline functions $\approx \Phi(L, W, S, B)$.
 - B-spline functions:
 - Fix m and consider

$$N_m(x_i) := \left(\underbrace{\mathbb{1}_{[0,1]} * \mathbb{1}_{[0,1]} * \cdots * \mathbb{1}_{[0,1]}}_{(m+1) \text{ times}} \right) (x_i).$$

- $N_m(x)$ is a piecewise polynomial of the order m .
- The following basis is called B-spline.

$$M_{k,j}^{m,d}(x) := \prod_{i=1}^d N_m(2^{k_i} x_i - j_i).$$

One can think of j as a location parameter (like 0th Haar wavelet basis) and k as spatial resolution (like k th Haar wavelet basis).

- $B_{p,q}^s(\Omega) \approx$ B-spline is established in (DP88).
- B-Spline $\approx \Phi(L, W, S, B)$ is from the following observations:
 - For some $M > 0$, write $\phi_{(0,M)}(x) := \sigma(x) - \sigma(x - M) = M \wedge \sigma(x)$.
 - Observe $N_m(x)$ has the form

$$N_m(x) = \frac{1}{m!} \sum_{j=0}^{m+1} (-1)^j \binom{m+1}{j} (m+1)^m \left(\phi_{(0,1-\frac{j}{m+1})} \left(\frac{x-j}{m+1} \right) \right)^m.$$

First, we focus on approximating $\left(\phi_{(0,1-\frac{j}{m+1})} \left(\frac{x-j}{m+1} \right) \right)^m$.

- (Yar17) showed for some $D \in \mathbb{N}$ there exists $\psi : \mathbb{R}^D \rightarrow \mathbb{R} \in \Phi(L_1, W_1, S_1, B_1)$ for some L_1, W_1, S_1, B_1 that depends on m and ϵ such that

$$\sup_{x \in [0, M]} \left| \psi \left(\underbrace{\left(\phi_{(0,M)} \left(\frac{x}{M} \right), \dots, \phi_{(0,M)} \left(\frac{x}{M} \right) \right)}_{m \text{ times. Write this function as } \psi \circ \phi_{(0,M)}(x/M)}. \right) - \left(\phi_{(0,M)} \left(\frac{x}{M} \right) \right)^m \right| \leq \epsilon$$

- Therefore, the reasonable construction of the approximator of $N_m(x)$ will be

$$f(x) = \frac{1}{m!} \sum_{j=0}^{m+1} (-1)^j \binom{m+1}{j} (m+1)^m \left(\psi \circ \phi_{(0,1-\frac{j}{m+1})} \left(\frac{x-j}{m+1} \right) \right)^m.$$

- Then, appropriately using ψ and f makes the form of $M_{0,0}^{m,d}(x)$.

- For any $F > 0$ and any function space $\mathcal{F} \subseteq B_{L^\infty(\Omega)}(0, F)$, there exists the following generalization gap type bound:

$$\mathbb{E}_{D_n} \left\| f^* - \hat{f} \right\|_{L^2(P)}^2 \leq C \left(\underbrace{\inf_{f \in \mathcal{F}} \|f^* - f\|_{L^2(P)}^2}_{\approx \|f^* - \tilde{f}\|^2} + \underbrace{(F^2 + \sigma^2) \frac{\log N(\mathcal{F}, \delta, \|\cdot\|_\infty)}{n}}_{\approx \mathbb{E} \|\hat{f} - \tilde{f}\|^2} + \delta(F + \sigma) \right).$$

Proof strategy:

- ① Substitute \hat{f} to the closest δ -minimal covering of \mathcal{F} and use the fact $\mathcal{F} \subseteq B_{L^\infty(\Omega)}(0, F)$ to bound the population risk by the empirical risk (Hardest part).
 - ② Bound the empirical risk in terms of the optimal recovery error: By using the fact that \hat{f} is ERM.
- Set $\mathcal{F} = \Phi(L, W, S, B) \cap B_{L^\infty(\Omega)}(0, F)$, and then the covering number analysis will give the following:

$$\log N(\Phi(L, W, S, B), \delta, \|\cdot\|_\infty) \leq 2SL \log((B \vee 1)(W + 1)) + S \log\left(\frac{L}{\delta}\right).$$

- Set $\delta = 1/n$, and in Step 1's RHS.

- Apply (1). the approximation result to get $\inf_{f \in \mathcal{F}} \|f^* - f\|_{L^2(P)}^2 \lesssim N^{-\frac{s}{d}}$, and (2). the covering number bound obtained in Step 2 with specific L, W, S, B in the approximation result.
- Then, optimizing the RHS w.r.t. N will induce the claimed bound with $N \asymp n^{\frac{d}{2s+d}}$.

- Since we are estimating the score (log-derivative) uniformly over the time t , there is a slight modification of the above result.
- Naively, this seems like a $d + 1$ dimensional and $s - 1$ smoothness function estimation problem. But, there is additional information for this problem: $P(X_t|X_0) \sim N(m_t X_0, \sigma_t^2)$ for some m_t, σ_t^2 .
- $\therefore P_t(x) = \int P_0(y) K_{\sigma_t^2}(\|x - m_t y\|^2) dy$ where $K_{\sigma_t^2}$ is a Gaussian kernel. Therefore, our target $\nabla \log P_t(x)$ also written as a fraction of $B_{p,q}^s * K_{\sigma_t^2}$.
- If we substitute $N_m(2^{k_i} x_i - j_i)$ in the B-spline by

$$E_{j,k}(x_i, t) = \int \mathbb{1}_{\{0,1\}}(2^{k_i} x_i - j_i) P_{N(m_t y_i, \sigma_t^2)}(x_i) dy_i,$$

Gaussian parts and Besov density parts separately controlled each other, and one can approximate $B_{p,q}^s * K_{\sigma_t^2}$ by $E_{j,k}(x_i, t)$. One can do the similar procedure as the above with this bases.

- \Rightarrow Population Score Loss of $\hat{s} \lesssim n^{-\frac{2s}{2s+d}} (\log n)^{16}$

- Using the Besov space estimation result, one can show

$$\sup_{P^*} \mathbb{E}_{D_n} TV(X_0, \hat{Y}_T) \lesssim n^{-\frac{s}{2s+d}} (\log n)^8.$$

- W_1 rate $n^{-\frac{s+1-\delta}{2s+d}}$ turned out to be faster. Why?
- Key observation: Utilizing the smoothness of the Gaussian noise.
 - Note the score network $s(X_t, t)$ does not have to be uniformly same over the time.
 - Observe $s_0 \approx \nabla B_{p,q}^s$, while $s_T \approx \nabla N(0, I)$.
 - Since $N(0, I)$ is very smooth, s_T is much easier to approximate/estimate than s_0 .
 - \therefore After the certain timestep t' , estimation error is expected to be much smaller.
 - Wrong but intuitive illustration:

$$d(\hat{P}_{[0,T]}, P_{[0,T]}^*) \leq \underbrace{d(\hat{P}_{[0,t']}, P_{[0,t']}^*)}_{\text{non-smooth target}} + \underbrace{d(\hat{P}_{[t',T]}, P_{[t',T]}^*)}_{\text{smooth target (Gaussian score)}}.$$

When $d = TV$, the 'non-smooth' term dominates, so cannot improve the DNN estimator rate. But when $d = W_1$, non-smooth part contributes less, so there is an improvement.

- ① For given $s, r \in [0, T]$, let $\bar{Y}^s(r)_t$ be a stochastic process s.t. $\bar{Y}^s(r)_0 = P_r$ and

$$d\bar{Y}^s(r)_t = \begin{cases} \beta_{T-t}(\bar{Y}^s(r)_t + 2\nabla \log P_{T-t}(\bar{Y}^s(r)_t))dt + \sqrt{2\beta_{T-t}}dB_t & t \in [0, T-s], \\ \beta_{T-t}(\bar{Y}^s(r)_t + 2\hat{s}_t(\bar{Y}^s(r)_t, T-t))dt + \sqrt{2\beta_{T-t}}dB_t & t \in [T-s, T]. \end{cases}$$

i.e., use the true score up to $T-s$ and then use the estimated score from $T-s$.

Particularly, one can think of $\bar{Y}^0(r)_t, \bar{Y}^T(r)_t$ similar to $Y_{T-r+t}, \hat{Y}_{T-r+t}$.

- ② Our target: $\mathbb{E}W_1(X_0, \hat{Y}_T) \leq \mathbb{E}W_1(\bar{Y}^T(T)_T, \hat{Y}_T) + \mathbb{E}W_1(X_0, \bar{Y}^T(T)_T)$.
- ③ First term: \hat{Y}_T and $\bar{Y}^T(T)_T$ only differs in the initial distributions ($N(0, I)$ and P_T resp.), leading to $\lesssim TV(N(0, I), P_T) \leq \exp(-\beta T)$ (\because reverse OU process).

④ Second term:

- Discretize $[0, T]$ by the partition made by $t_j = C^j n^{-\frac{2-\delta}{2s+d}}$ for some $j = 1, \dots, k = O(\log n)$. Here C is a constant that makes $C^k n^{-\frac{2-\delta}{2s+d}} = T$. A certain t_j will be t' mentioned above.
- Important: This interval is not 'equi-length'. Smaller t has the smaller interval.
- $\mathbb{E} W_1(X_0, \bar{Y}^T(T)_T) \leq \sum_j \mathbb{E} W_1(\bar{Y}^{j-1}(T)_T, \bar{Y}^j(T)_T)$.
- $\bar{Y}^{j-1}(T)_T$ and $\bar{Y}^j(T)_T$ has the same initial distribution as well as the dynamics, except the difference in the drift term of $[t_{j-1}, t_j]$.
- Girsanov Theorem gives the KL bound of such processes in terms of the difference between drift terms, and (omitting the complicated steps) leads to

$$\mathbb{E}_{x_0} W_1(\bar{Y}^{j-1}(T)_T, \bar{Y}^j(T)_T) \lesssim \sqrt{t_j \log n \int_{t_{j-1}}^{t_j} \mathbb{E}_{x_t, x_0 \sim P_t, P_0} \|\hat{S}(x_t, t) - \nabla \log P_t(x_t)\|^2 dt} + n^{-\frac{s+1}{2s+d}}.$$

Note the bound gets smaller when t_j is small (corresponding to the Key Observation).

- Plug-in the estimation error bound of the score loss (in the interval $[t_{j-1}, t_j]$), and plug-in $t_j = C^j n^{-\frac{2-\delta}{2s+d}}$ to derive the desired value.

- The extra term δ appears in W_1 from optimizing the choice of the threshold t' .
- In case of TV distance, one obtains the bound as in the above, but with t_j part substituted by $O(1)$. So, one cannot tighten the bound when t_j is small, so the error of non-smooth part equally contributes.
- One can think of this as 'Mean-diff ($\approx W_1$) \leq Max-diff (\approx TV)' type inequality.

- DNN diffusion model with score matching loss achieves almost minimax rate w.r.t. W_1 (and TV) distance.
- The fundamental ingredient is from the minimax function estimation in Besov space.
 - Approximation theory to obtain the good approximator (B-Spline in Besov case).
 - Learning theory to bound the gap between estimator and the approximator.
- The OU process structure of the diffusion model gives some advantage:
 - Target score has the same smoothness as the density.
 - As $t \rightarrow T$, target score gets smoother.

- How to actually train such a constraint neural network?
 - Can we reformulate the constraint into tractable ways, e.g., unconstrained optimization with appropriate regularizer?
 - Constraints play a role in two parts:
 - ① Approximation: S, B enables to avoid the overfitting to the noise (\approx LASSO type regularizer).
 - ② Learning: S, B enables to bound the covering number, which controls the generalization bound.
 - It is not immediate how to avoid such constraints in the approximation stage: Relationship to weight decay type penalty?
 - On the other hand, there are alternative approaches to obtain a generalization bound (e.g., Rademacher complexity) to avoid the constraint. Can we utilize those?
 - PAC (Bayes) type analysis for the specific algorithm?
- Adaptivity
 - Constructing $\Phi(L, W, S, B)$ requires the prior knowledge on the regularity of the P^* ; e.g., choices of L, W, S, B require s, p . This makes the estimation non-adaptive.
- Claims using ‘two’ DNNs at $[0, t']$ and $[t', T]$ improves the rate. When to exactly? Can $\hat{s}(x_t, t)$ be adaptive to t' ?

Thank You For Your Attention!

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